

# Introduction to “Low-Degree Method” for Computational Hardness

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(based on a survey with Afonso Bandeira and Alex Wein,  
**which is based on deep ideas not original to us!**)

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## Question:

How to predict when **statistical inference**  
will be **computationally hard**?

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For this talk, **statistical inference = hypothesis testing.**

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- You see  $Y$ , and try to *infer* which one using a *test*:

$$f : \mathbb{R}^N \rightarrow \{p, q\}$$

# What is **asymptotic** statistical inference?

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This lets us define **asymptotic success** (“strong detection”):

$$\lim_{n \rightarrow \infty} \mathbb{P}_n[f_n(\mathbf{Y}) = p] = 1,$$

$$\lim_{n \rightarrow \infty} \mathbb{Q}_n[f_n(\mathbf{Y}) = q] = 1.$$

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  - $\mathbb{P}_n: (\mathbf{g}_1, \dots, \mathbf{g}_{kn}) \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_n + \lambda \mathbf{x} \mathbf{x}^\top)$
- Community detection
  - $\mathbb{Q}_n: G \sim \text{Erdős-Rényi}$
  - $\mathbb{P}_n: G \sim \text{Erdős-Rényi} + \text{clique}$   
 $G \sim \text{different edge prob. within/between blocks}$

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- Spiked transport model [Rigollet, Weed 2019]<sup>1</sup>
  - $\mathbb{Q}_n$ :  $(\mathbf{x}_1, \dots, \mathbf{x}_m), (\mathbf{y}_1, \dots, \mathbf{y}_m)$  i.i.d.
  - $\mathbb{P}_n$ :  $\mathbf{x}_i = \mathbf{a}_i^{(1)} + \mathbf{z}_i^{(1)}, \mathbf{y}_i = \mathbf{a}_i^{(2)} + \mathbf{z}_i^{(2)}$   
 $\mathbf{a}^{(j)}$  different laws on low-dimensional subspace  $V$ , and  
 $\mathbf{z}^{(j)}$  same law on  $V^\perp$ .

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Optimizer: the (normalized) **likelihood ratio**

$$h^*(\mathbf{Y}) = \frac{d\mathbb{P}}{d\mathbb{Q}}(\mathbf{Y}) \ / \ \underbrace{\left\| \frac{d\mathbb{P}}{d\mathbb{Q}} \right\|}_{\text{objective value}}$$

# Justification 1: optimal error tradeoff

[Neyman, Pearson 1933] Of tests with  $\mathbb{Q}[f(\mathbf{Y}) = p] \leq \alpha$ , the test that minimizes  $\mathbb{P}[f(\mathbf{Y}) = q]$  is

$$f_{\xi}(\mathbf{Y}) = \left\{ \begin{array}{ll} p & \text{if } \frac{d\mathbb{P}}{d\mathbb{Q}}(\mathbf{Y}) \geq \xi \\ q & \text{otherwise.} \end{array} \right\},$$

for suitable  $\xi$ .

**Best tradeoff between “Type I” and “Type II” errors.**

**(And non-asymptotically!)**

## Justification 2: control of **asymptotic** success

[Le Cam, 1960's] Suppose  $\| \frac{d\mathbb{P}_n}{d\mathbb{Q}_n} \| \leq K$  as  $n \rightarrow \infty$ . Then,  $\mathbb{P}_n$  is *contiguous* to  $\mathbb{Q}_n$ :

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**Corollary:** Set  $A_n = \{f_n(\mathbf{Y}) = p\}$ . Then:

$$\underbrace{\mathbb{Q}_n[f_n(\mathbf{Y}) = p] \rightarrow 0}_{\text{success under } \mathbb{Q}_n} \Rightarrow \underbrace{\mathbb{P}_n[f_n(\mathbf{Y}) = p] \rightarrow 0}_{\text{failure under } \mathbb{P}_n}.$$

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“Information-theoretic” (no efficiency worries) limitations:

$$\| \frac{d\mathbb{P}_n}{d\mathbb{Q}_n} \| \text{ bounded} \Rightarrow \text{no test succeeds}$$

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 $p \in \mathbb{R}[y_1, \dots, y_N]$  with  $\deg(p) \leq D$  computable in  $O(N^D)$ .

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Optimizer: the (normalized) **low-degree likelihood ratio**

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# The low-degree conjecture

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**Main conjecture:**

$$\| \mathcal{P}^{\leq (\log N)^{1+\epsilon}} \frac{d\mathbb{P}_n}{d\mathbb{Q}_n} \| \text{ bounded} \Rightarrow \text{no efficient test succeeds}$$



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One wrinkle: rather than  $D = \omega(1)$ , to include calculation of spectral norms of matrices  $\rightsquigarrow D = \omega(\log N)$ .

**Main conjecture:**

$$\| \mathcal{P}^{\leq (\log N)^{1+\epsilon}} \frac{d\mathbb{P}_n}{d\mathbb{Q}_n} \| \text{ bounded} \Rightarrow \text{no efficient test succeeds}$$

Originally from sum-of-squares optimization (fancy semidefinite programming) literature: controls whether a lower bound construction succeeds or not.

- [Barak, Hopkins, Kelner, Kothari, Moitra, Potechin 2016]
- [Hopkins, Steurer 2017]
- [Hopkins, Kothari, Potechin, Raghavendra, Schramm, Steurer 2017]
- [Hopkins 2018] (PhD thesis)

$$\limsup_{n \rightarrow \infty} \left\| \mathcal{P}^{\leq (\log N)^{1+\epsilon}} \frac{d\mathbb{P}_n}{d\mathbb{Q}_n} \right\| = \begin{cases} +\infty & \rightsquigarrow \text{maybe easy} \\ K & \rightsquigarrow \text{hard} \end{cases}$$

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### Question 1:

How to project to low-degree polynomials?

### Question 2:

How to evaluate asymptotics?

# A simple gaussian model

Let's restrict to a special case to show how this works:

- $\mathcal{P}_n$  a “prior” over  $\mathbb{R}^N$ .
- $\mathcal{Q}_n$ :  $\mathbf{Y} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_N)$ .
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A very special case:  $N(n) = n^2$ ,  $\mathcal{P}_n$  distribution over rank 1 matrices  $\mathbf{X} = \sqrt{\frac{n}{2}}\lambda\mathbf{x}\mathbf{x}^\top$ , e.g.,  $\mathbf{x} \sim \text{Unif}(\mathbb{S}^{n-1})$ . Symmetrizing,

$$\underbrace{\text{GOE}(n)}_{\mathcal{Q}_n} \quad \text{vs.} \quad \underbrace{\text{GOE}(n) + \sqrt{n} \cdot \lambda\mathbf{x}\mathbf{x}^\top}_{\mathbb{P}_n}$$

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[Féral, Pécché 2007] Top eigenvalue test succeeds iff  $\lambda > 1$ .

**Question:** Is this optimal?

# Step 1: computing the likelihood ratio

The model:

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For likelihood ratio, just need gaussian densities:

$$\frac{d\mathbb{P}_n}{d\mathbb{Q}_n}(\mathbf{Y}) = \mathbb{E}_{\mathbf{X} \sim \mathcal{P}_n} \left[ \frac{d\mathbb{P}_n[\cdot | \mathbf{X}]}{d\mathbb{Q}_n}(\mathbf{Y}) \right]$$



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## Step 2: computing the low-degree projections

Use the orthogonal basis of **Hermite polynomials**,

$$h_k(y) \in \mathbb{R}[y]$$

$$H_k(\mathbf{Y}) = \prod_{i=1}^N h_{k_i}(Y_i) \in \mathbb{R}[Y_1, \dots, Y_N]$$

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Projections by **generalized gaussian integration by parts**:

$$\left\langle \frac{d\mathbb{P}_n}{d\mathbb{Q}_n}, H_k \right\rangle = \mathbb{E}_{\mathbf{Y} \sim \mathbb{Q}_n} \left[ \frac{\partial^{\sum k_i}}{\partial Y_1^{k_1} \dots \partial Y_N^{k_N}} \frac{d\mathbb{P}_n}{d\mathbb{Q}_n} \right]$$

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## Step 3: computing the norm

Use (part of) the **replica trick** to handle squared  $\mathbb{E}[\dots]$ :

$$\begin{aligned}\left\| \mathbf{p}_{\leq D} \frac{d\mathbb{P}_n}{d\mathbb{Q}_n} \right\|^2 &= \sum_{\sum k_i \leq D} \frac{1}{\prod k_i!} \left\langle \frac{d\mathbb{P}_n}{d\mathbb{Q}_n}, H_{\mathbf{k}} \right\rangle^2 \\ &= \sum_{\sum k_i \leq D} \frac{1}{\prod k_i!} \left( \mathbb{E}_{\mathbf{x} \sim \mathbb{P}_n} [\prod x_i^{k_i}] \right)^2\end{aligned}$$

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## Step 4: evaluating the asymptotic

The special case:  $\mathbf{X} = \sqrt{n/2} \cdot \lambda \mathbf{x} \mathbf{x}^\top$ ,  $\mathbf{x} \sim \text{Unif}(\mathbb{S}^{n-1})$ .

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**Key:**  $D(n) \ll n$ , so CLT “kicks in” in time for moments.



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$$\underbrace{\mathbb{E}_{g \sim \mathcal{N}(0,1)} \exp \left( \frac{\lambda^2}{2} g^2 \right)}_{\text{natural, scalar expectation!}}$$

# Review

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# Review

1. The *low-degree conjecture* connects hardness of statistical testing with the norm of the *low-degree likelihood ratio*.
2. To analyze a problem, we proceed as follows:
  - 2.1 Compute the likelihood ratio
  - 2.2 Find the orthogonal polynomials of the null model ( $\mathbb{Q}$ )
  - 2.3 Project (using **special distributional properties**)
  - 2.4 Compute the norm (using “**baby replica trick**”)
  - 2.5 Reduce to scalar expectation (**limit theorem** heuristic)

# Other frameworks for hardness predictions

1. Conjecturally optimal algorithms
  - 1.1 BP / AMP  $\sim$  cavity and replica methods of stat. physics
  - 1.2 Sum-of-squares hierarchy (semidefinite programming)
  - 1.3 Monte Carlo sampling from posterior
  - 1.4 Local algorithms
  - 1.5 Problem-specific algorithms (e.g. PCA)
2. Structure of solution space (“shattering” & co.)
3. Geometric analysis of optimization landscapes
4. Average-case reductions

# The bright side

The low degree method is...

- Easy
- Uniform across problems
- Broadly applicable (to nice “toy-ish” setups)
- Intuitively plausible
- Always correct (so far)



# The other hand

The low degree method is...

- Coarse-grained in runtimes
- Hard to handle correlated models with
- **Dependent on orthogonal polynomial magic**
- Dependent on good control of signal priors
- Not a great way to design actual algorithms

So...give it a try when you are wearing your theorist hat, and want to make a **quick, painless prediction of thresholds for a nice model.**

**Thank you!**